

Estadística bayesiana y aplicaciones en ciencia de datos

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1 Statistical induction

2 Dirichlet process: the canonical BNP prior

3 BNP mixtures

- 1 Statistical induction
- 2 Dirichlet process: the canonical BNP prior
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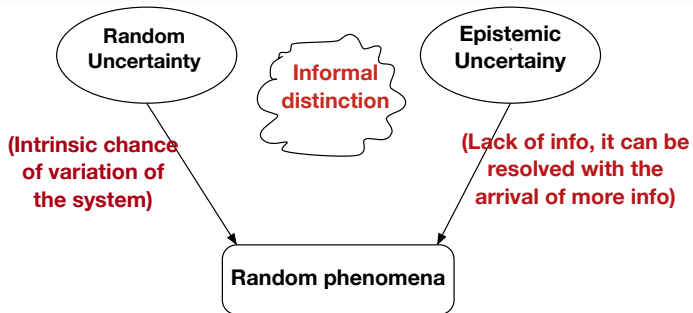
Random phenomena drive many aspects of this world



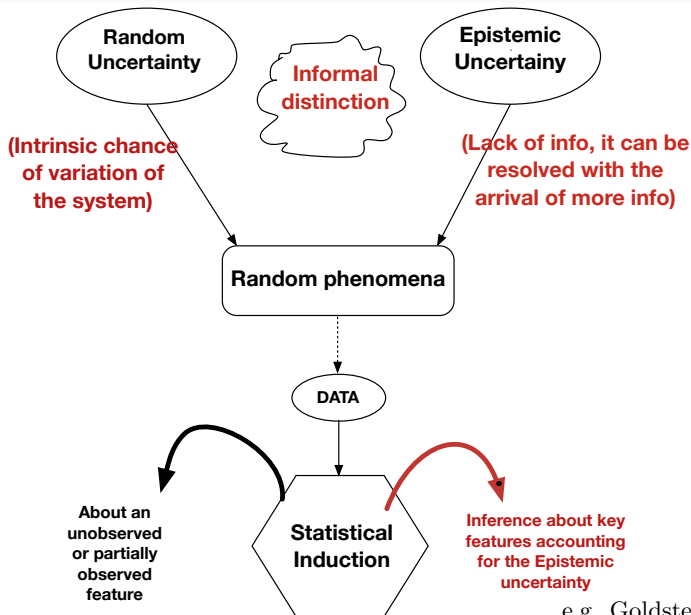
The Bayesian approach to statistical induction

Random phenomena

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The Bayesian approach to statistical induction



The basic probabilistic setup

- $(\Omega, \mathcal{A}, \mathbb{P})$: *Probability space*
 - ▷ Ω –*sample space*. Set of all possible outcomes
 - ▷ \mathcal{A} – *σ -field*. Collection of subsets of Ω with all events of interest
 - ▷ $\mathbb{P} : \mathcal{A} \mapsto [0, 1]$ –*Probability measure*. Mathematically coherent measure to quantify all *events* $A \in \mathcal{A}$
- Features of interest can be translated into “numeric” quantities via
 - ▷ $(\mathbb{X}, \mathcal{X})$ -valued functions, $X : \Omega \mapsto \mathbb{X}$. *random variables* (r.v.’s)
- Given a r.v. X , the set function defined by

$$P_X(B) = \mathbb{P}(X^{-1}(B)), \quad \text{for all } B \in \mathcal{X} \quad (1)$$

is termed the *distribution or law* of the random variable X .

- ▷ When $\mathbb{X} = \mathbb{R}$ and $B = (-\infty, x]$ we write

$$F_X(x) = P_X((-\infty, x]) = \mathbb{P}(X \leq x) \quad \rightarrow \text{the (cdf) of } X$$

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Ex. toss a coin

$$\Omega = \{\text{head, tail}\} = \{\omega_1, \omega_2\} = \{0, 1\}$$

$$\mathcal{A} = \{\Omega, \{0\}, \{1\}, \emptyset\}$$

Let X the r.v. that assigns 1 if the outcome is tail and 0 otherwise, i.e. $P_X(\{1\}) = \mathbb{P}(X(\omega_1) = 1)$ with $\mathbb{X} = \{0, 1\}$

- For such quantity, we might assign a value $\theta \in [0, 1]$, i.e.

$$P_X(\{1\}) = \theta$$

⇒ Uncertainty about X is transferred to the *parameter* of interest θ .

How can we improve our knowledge about θ in the presence of observations from the random phenomena?

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- Availability of more info about a random phenomenon
 - ⇒ better uncertainty quantification
 - ⇒ better statistical induction
- Realizations of a given phenomenon encoded via r.v.'s $\{X_i\}_{i \in \mathcal{I}}$
 - ▷ Logical/physical independence $\not\Rightarrow$ stochastic independence
 - so $\mathbb{P}(X_{n+1} \in B \mid X_1, \dots, X_n) = \mathbb{P}(X_{n+1} \in B)$ not always a good idea!
 - ▷ Statistical learning requires stochastic dependence !

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- Symmetry/Stability principles in the law modelling $\{X_i\}$'s are fundamental for statistical induction
 - ▷ e.g. the past and future have similar behaviour
- Major symmetries used in statistics
 - ▷ IID r.v.'s: physical & stochastic independence (rare in real apps!)
 - ▷ Exchangeability: physical indep. + sampling order invariance!
 - ▷ Stationarity: Uncertainty is not “time” invariant
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Exchangeable sequences

A finite sequence of r.v.'s, $\{X_n\}_{i=1}^n$, is said to be *finite exchangeable* if, for any permutation π of $(1, \dots, n)$

$$(X_1, \dots, X_n) \stackrel{d}{=} (X_{\pi(1)}, \dots, X_{\pi(n)})$$

An infinite sequence $\{X_n\}_{i=1}^{\infty}$ is said to be *exchangeable* if every subcollection is exchangeable.

≈ Distributional invariance under sampling order

▷ What can we say about the law of an exchangeable sequence

▷ B. de Finetti's representation characterises exchangeable sequences

de Finetti's representation Theorem: $\mathbb{X} = \{0, 1\}$ case

- B. de Finetti 1931: A seq. of binary r.v.'s $\{X_i\}_{i=1}^{\infty}$, e.g. with values in $\mathbb{X} = \{0, 1\}$, is exchangeable iff there exists a dist. \mathbf{q} on $[0, 1]$

$$\mathbb{P}(X_1 = x_1, \dots, X_n = x_n) = \int_{[0,1]} \theta^{s_n} (1 - \theta)^{n - s_n} \mathbf{q}(d\theta)$$

where $s_n := \sum_{i=1}^n x_i$.

- $\mathbf{q}(\cdot)$ is the distribution of $\lim_{n \rightarrow \infty} \frac{s_n}{n}$
- Conditional independence

$$\mathbb{P}(X_1 = x_1, \dots, X_n = x_n \mid \theta) = \prod_{i=1}^n \mathbb{P}(X_i = x_i \mid \theta) = \theta^{s_n} (1 - \theta)^{n - s_n}$$

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Example: $\{X_i\}_{i=1}^{\infty}$ be Bernoulli r.v.'s

Two different Bernoulli exchangeable laws by two different persons

$$\mathbb{P}(x_1, \dots, x_n) = \frac{12}{s_n + 2} \frac{1}{\binom{n+4}{s_n+2}} \quad \text{and} \quad \mathbb{P}(x_1, \dots, x_n) = \frac{1}{[n+1] \binom{n}{s_n}},$$

- ▷ These persons believe that $\mathbb{P}(X_1 = 1) = 0.4$ & $\mathbb{P}(X_1 = 1) = 0.5$ resp.
- ▷ Both believe that $\Theta := \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i$ exists & $\mathbb{P}(X_1 = 1 \mid \Theta = \theta) = \theta$
- ▷ Since $\mathbb{P}(X_1 = 1) = \mathbb{E}(\Theta)$ they must have different values for $\mathbb{E}(\Theta)$
- ▷ Assume we observe the result of $n = 20$ given by 14 “1s” and 6 “0s”.

$$\mathbb{P}[X_{21} \mid x_1, \dots, x_{20}] = 0.64 \quad \text{and} \quad \mathbb{P}[X_{21} \mid x_1, \dots, x_{20}] = 0.68.$$

- Regardless of the prior mean on Θ , they should modify their opinion about the prop. of 1's!
- Consequence due to exchangeability, regardless of frequencies being interpreted as probabilities.

In Bayesian terms

If X is a r.v. on $\mathbb{X} = \{1, 2\}$ with prob. p_1 and $p_2 = 1 - p_1$ assigned to each element of \mathbb{X} . That is $\{X \mid p_1, p_2\} \sim \text{Bernoulli}(p_1, p_2)$

- Each value of $\mathbf{p} = (p_1, p_2)$ defines probability measure on \mathbb{X}
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$\mathcal{P}_{\mathbb{X}} := \{\text{Space of prob. measures on } \mathbb{X}\} = \{(p_1, p_2); p_i \geq 0 \text{ y } p_1 + p_2 = 1\}$

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One parameter per value in the support \mathbb{X} !

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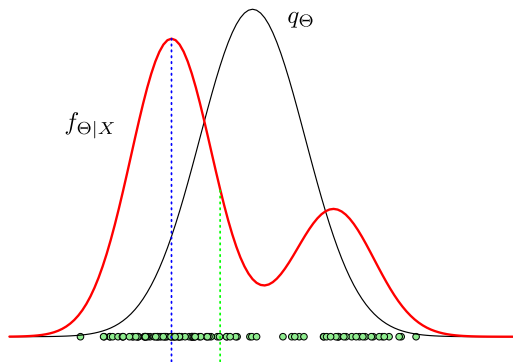
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Prior to posterior effect



Exchangeable sequences: general \mathbb{X}

- Let $\mathcal{P}_{\mathbb{X}}$ be the space of all probability measures on $(\mathbb{X}, \mathcal{X})$

A seq. $\{X_i\}_{i=1}^{\infty}$ is **exchangeable** iff there exists \mathbf{Q} on $\mathcal{P}_{\mathbb{X}}$ such that

$$\mathbb{P}(X_1 \in A_1, \dots, X_n \in A_n) = \int_{\mathcal{P}_{\mathbb{X}}} \prod_{i=1}^n \mathbb{P}(A_i) \mathbf{Q}(d\mathbb{P}), \quad \forall n \geq 1 \text{ and } A_i \in \mathcal{X}$$

Alternatively: $X_i \mid \mathbb{P} \stackrel{\text{iid}}{\sim} \mathbb{P}$ and $\mathbb{P} \sim \mathbf{Q}$ (**conditionally iid**).

Hewitt and Savage 1955

- ▷ If $P_n(A) := \frac{1}{n} \sum_{i=1}^n \delta_{X_i}(A)$ denotes the **empirical dist.** hence, \mathbf{Q} is the dist. of the RPM \mathbf{P} , where $\mathbb{P}[P_n \Rightarrow \mathbf{P}] = 1$ ($\mathbf{P} \sim \mathbf{Q}$)
- ▷ \mathbf{Q} is unique
- ▷ “The unknown”, \mathbf{P} , that allows us to disaggregate the elements of $X^{(\infty)}$ as a conditional iid sample, **is random**.

Consequences of de Finetti's representation

- There is a clear **bijection** between the law of $\{X_i\}_{i=1}^{\infty}$ and \mathbb{Q}
 - Pick $\mathbb{Q} \Rightarrow$ we have a law for $\{X_i\}_{i=1}^{\infty}$
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 i.e. $\mathbb{P}[X_{n+1} \in A \mid X^{(n)}] = \mathbb{E}_{\mathbb{Q}_{X^{(n)}}} [\mathbb{P}(A)]$ characterizes \mathbb{Q} with
 $\mathbb{Q}_{X^{(n)}}(B) := \mathbb{P}(P \in B \mid X^{(n)})$,
- However, any $\mathbb{P} \in \mathcal{P}_{\mathcal{X}}$ can be seen as the limit of P_n !
- Bayesian interpretation:
 \mathbb{Q} takes the interpretation of prior distributions on \mathcal{P}

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i.e. $\mathbb{P}[X_{n+1} \in A \mid X^{(n)}] = \mathbb{E}_{\mathbb{Q}_{X^{(n)}}} [\mathbb{P}(A)]$ characterizes \mathbb{Q} with
 $\mathbb{Q}_{X^{(n)}}(B) := \mathbb{P}(\mathbb{P} \in B \mid X^{(n)})$,
- However, **any** $\mathbb{P} \in \mathcal{P}_{\mathbb{X}}$ can be seen as the limit of P_n !
- Bayesian interpretation:
 \mathbb{Q} takes the interpretation of prior distributions on \mathbb{P}

Probabilistically speaking the Bayesian approach is equivalent to the exchangeability assumption of the $\{X_i\}_{i=1}^{\infty}$

de Finetti and the Bayesian approach

The law of the exchangeable r.v.'s (and thus Q) is characterized by the conditional probabilities (or **predictive distributions**)

$$\begin{aligned} \mathbb{P}[X_{n+1} \in A_{n+1} \mid X_1 \in A_1, \dots, X_n \in A_n] &= \frac{E_Q \left[\prod_{i=1}^{n+1} P(A_i) \right]}{E_Q \left[\prod_{i=1}^n P(A_i) \right]} \\ &= E_{Q_{X^{(n)}}} [P(A_{n+1})] \end{aligned}$$

for all $n > 1$, with $P_0 := \mathbb{P}[X_1 \in A_1] = E_Q[P(A_1)]$ and where

$$Q_{X^{(n)}}(dP) = \frac{\prod_{i=1}^n P(A_i) Q(dP)}{E_Q \left[\prod_{i=1}^n P(A_i) \right]}, \quad (\text{dominated case})$$

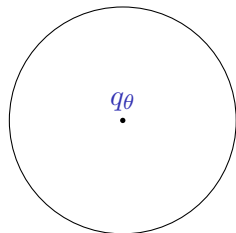
the **posterior distribution** of P given $X^{(n)} := (X_1, \dots, X_n)$

Exchangeability: statistical learning for physically independent observations

Random phenomena encoded in \mathbb{X} -valued $\{X_i\}_{i=1}^{\infty}$ exchangeable sequence driven by $\mathbf{P} \sim \mathbf{Q}$

- $\mathbf{Q}(\cdot) = \delta_{q_\theta}(\cdot) \Rightarrow X_i$'s are iid

$$\mathbb{P}(X_1 \in A_1, \dots, X_n \in A_n) = \int_{\mathcal{P}_{\mathbb{X}}} \prod_{i=1}^n \mathbb{P}(A_i) \delta_{q_\theta}(d\mathbb{P}) = \prod_{i=1}^n q_\theta(A_i)$$



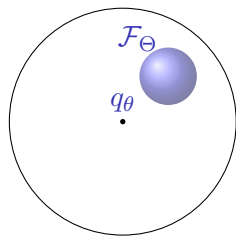
$\mathcal{P}_{\mathbb{X}}$: Space of all distributions on \mathbb{X}

Exchangeability: statistical learning for physically independent observations

Random phenomena encoded in \mathbb{X} -valued $\{X_i\}_{i=1}^{\infty}$ exchangeable sequence driven by $\mathbf{P} \sim \mathbf{Q}$

- $\mathbf{Q}(\mathcal{F}_{\Theta}) = 1 \Rightarrow$ Parametric family

$$\mathbb{P}(X_1 \in A_1, \dots, X_n \in A_n) = \int_{\mathcal{F}_{\Theta}} \prod_{i=1}^n \underbrace{F_{\theta}(A_i)}_{\text{Random uncertainty via param. model}} \overbrace{\pi_{\theta}(d\theta)}^{\text{Epistemic uncertainty}}$$



$\mathcal{P}_{\mathbb{X}}$: Space of all distributions on \mathbb{X}

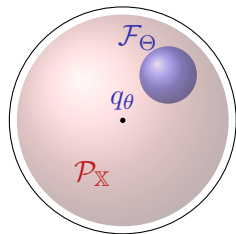
Exchangeability: statistical learning for physically independent observations

Random phenomena encoded in \mathbb{X} -valued $\{X_i\}_{i=1}^{\infty}$ exchangeable sequence driven by $P \sim Q$

- $Q(P : d(P, \eta) < \varepsilon) > 0, \forall \eta \in \mathcal{P}_{\mathbb{X}} \text{ y } \varepsilon > 0 \Rightarrow \text{BNP}$

$$\mathbb{P}(X_1 \in A_1, \dots, X_n \in A_n) = \int_{\mathcal{P}_{\mathbb{X}}} \prod_{i=1}^n \underbrace{P(A_i) Q(dP)}$$

Random and
epistemic uncertainties
in one stroke!



$\mathcal{P}_{\mathbb{X}}$: Space of all distributions on \mathbb{X}
...or other infinite dimensional
sub-spaces of interest, $\mathcal{P}_{\mathbb{X}}^d, \mathcal{P}_{\mathbb{X}}^c$, etc.

Bayesian nonparametrics

What happens if \mathbb{X} is of an infinite nature?

- ▷ We could $\mathcal{P}_{\mathbb{X}}|_{\mathcal{F}_{\Theta}}$, but doesn't resolve the "random uncertainty"
- ▷ We want models Q giving positive prob. to all elements of $\mathcal{P}_{\mathbb{X}}$, or at least some infinite subset, e.g. set of densities, cdf's, etc.
- ▷ de Finetti's representation Th. for general \mathbb{X} gives an answer...
- ▷ Remember: $\{X_i\}_{i=n}^{\infty}$ exchangeable is driven by $P \sim Q$

How to construct suitable models for Q (nonparametric priors!)?

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How to construct suitable models for Q (nonparametric priors!)?

The Dirichlet distribution

Let $Z_i \stackrel{\text{iid}}{\sim} \text{Ga}(\alpha_i, 1)$, $i = 1, \dots, m$ and $\mathbf{W} := (W_1, \dots, W_m)$ with

$$W_i = \frac{Z_i}{\sum_{i=1}^m Z_i}, \quad i = 1, \dots, m \quad \Rightarrow \quad \mathbf{W} \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_m)$$

and is independent of $Z := \sum_{i=1}^m Z_i \sim \text{Ga}(\sum_{i=1}^m \alpha_i, 1)$ with density

$$f(\mathbf{w}) = \frac{\Gamma(\sum_{i=1}^m \alpha_i)}{\prod_{i=1}^m \Gamma(\alpha_i)} \prod_{i=1}^{m-1} w_i^{\alpha_i-1} \left(1 - \sum_{i=1}^{m-1} w_i\right)^{\alpha_m-1} \mathbb{I}_{\Delta_{m-1}}(\mathbf{w}),$$

where $\Delta_{m-1} := \left\{ (w_1, \dots, w_{m-1}) : w_i \geq 0, \sum_{i=1}^{m-1} w_i \leq 1 \right\}$

Properties of Dirichlet distribution

Moments

Let $\alpha := \sum_{i=1}^m \alpha_i$ and $p_i := \alpha_i/\alpha$ hence

- $E[w_i] = p_i$
- $\text{Var}[w_i] = \frac{p_i(1-p_i)}{\alpha+1}$
- $\text{Corr}[w_i, w_j] = -\frac{p_i p_j}{\alpha+1}$

Addition property

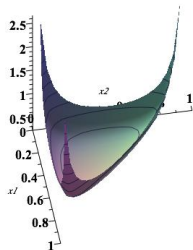
If $W \sim \text{Dirichlet}(\alpha_1, \alpha_2, \dots, \alpha_m)$ then

- i) For any partition A_1, \dots, A_k of $\{1, \dots, n\}$, the vector

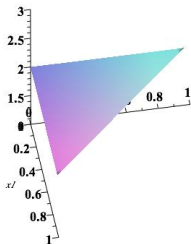
$$\left(\sum_{i \in A_1} w_i, \sum_{i \in A_2} w_i, \dots, \sum_{i \in A_k} w_i \right) \sim \text{Dirichlet}(\alpha'_1, \dots, \alpha'_k)$$

where $\alpha'_i := \sum_{j \in A_i} \alpha_j$

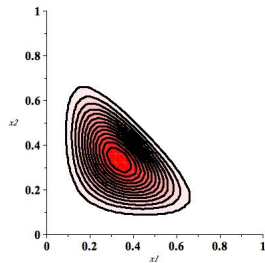
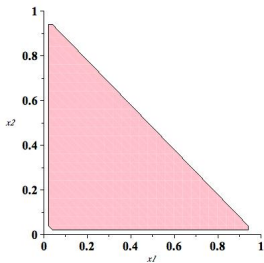
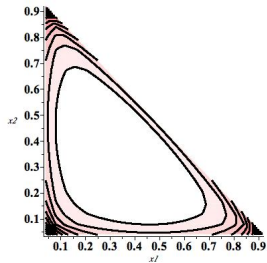
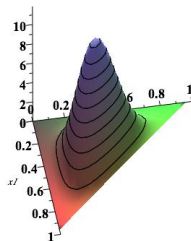
$$\alpha_1 = \alpha_2 = \alpha_3 = 0.2$$



$$\alpha_1 = \alpha_2 = \alpha_3 = 1$$



$$\alpha_1 = \alpha_2 = \alpha_3 = 5$$



Ferguson 1973: The canonical example

1 Via infinite dimensional distributions with pre-scribed fdds

Let $\alpha > 0$ a non-atomic finite measure on a Polish space $(\mathbb{X}, \mathcal{X})$. A $\mathcal{P}_{\mathbb{X}}$ -valued RPM, \mathbf{P} , is said to have a **Dirichlet process** (\mathcal{D}_{α}) distribution, if for all measurable partition (B_1, \dots, B_k) de \mathbb{X}

$$(\mathbf{P}(B_1), \dots, \mathbf{P}(B_k)) \sim \text{Dir}(\alpha(B_1), \dots, \alpha(B_k))$$

- Ferguson 73' proved that the Dirichlet dist. is projective and therefore Daniel-Kolmogorov's existence theorem ensures the existence of \mathcal{D}_{α} . Namely, a stochastic process indexed on \mathcal{X} .

The Dirichlet process \mathcal{D}_α : The canonical example

Extending the finite-dim properties to the infinite-dim object it can be seen that if $X_i \stackrel{\text{iid}}{\sim} P$ and $P \sim \mathcal{D}_\alpha$ then

- $P_0(B) := E_{\mathcal{D}_\alpha}[P] = \frac{\alpha(B)}{\theta}$ for $B \in \mathcal{X}$ and where $\theta := \alpha(\mathbb{X})$
- $\text{Var}_{\mathcal{D}_\alpha}[P(B)] = \frac{P_0(B)(1-P_0(B))}{\theta+1}$
- $\text{Cov}(P(B_1), P(B_2)) = \frac{P_0(B_1 \cap B_2) - P_0(B_1)P_0(B_2)}{\theta+1}$

If $X_i | P \stackrel{\text{iid}}{\sim} P$ y $P \sim \mathcal{D}_{\theta P_0}$, then $X_i \sim P_0, \forall i = 1, 2, \dots$

$$P | X_1, \dots, X_n \sim \mathcal{D}_{\theta P_0 + n P_n} \quad (\text{conjugacy})$$

$$E[P | X_1, \dots, X_n] = \frac{\theta}{\theta + n} P_0 + \frac{n}{\theta + n} \sum_{i=1}^n \frac{\delta_{X_i}}{n}, \quad (\text{Bayes estimator})$$

The Dirichlet process \mathcal{D}_α : Pólya urn representation

2 Specification of Q via predictive distributions.

- Q can be characterized by its predictive dist. (Bayes estimator)

$$\mathbb{P}(X_{n+1} \in A \mid X_1, \dots, X_n) = \mathbb{E}[P(A) \mid X_1, \dots, X_n] = \frac{\alpha_n(A)}{\alpha_n(\mathcal{X})}$$

with $\alpha_n(\cdot) = \alpha(\cdot) + \sum_{i=1}^n \delta_{X_i}(A)$. In other terms

$$\mathbb{P}[X_{n+1} \in \cdot \mid X^{(n)}] = \underbrace{\frac{\theta}{\theta + n}}_{\mathbb{P}[X_{n+1} = \text{"new"} \mid X^{(n)}]} \underbrace{P_0(\cdot)}_{\text{Prior guess}} + \underbrace{\frac{n}{\theta + n}}_{\mathbb{P}[X_{n+1} = \text{"old"} \mid X^{(n)}]} \underbrace{\sum_{i=1}^n \frac{\delta_{X_i}(\cdot)}{n}}_{\text{empirical measure}},$$

Q is a DP iff the predictive is a linear combination of P_0 and the empirical measure

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The Dirichlet process \mathcal{D}_α : Pólya urn representation

Blackwell and MacQueen '73' observed that when $n \rightarrow \infty$

$$\frac{\alpha_n(\cdot)}{\alpha_n(\mathbb{X})} \xrightarrow{\text{a.s.}} \mathbf{P}, \quad \text{with} \quad \mathbf{P} \sim \mathcal{D}_\alpha$$

→ Very appealing for MCMC implementations

→ A direct consequence is that

$$\mathbb{P}(X_i = X_j) = \frac{1}{\theta + 1} > 0, \quad i \neq j$$

Blackwell '73' proved that

$$\mathbb{P}(\mathbf{P} \text{ is discrete}) = 1$$

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